MATH 60604A Statistical modelling § 5h - Group heteroscedasticity

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- Assume that observations within groups have the same covariance structure, but the parameters of the latter differ between groups.
- Assuming consecutive grouped measurements, the covariance matrix of all the measurements is

$$\operatorname{Cov}(Y) = \begin{pmatrix} \boldsymbol{\Sigma}_1 & \boldsymbol{0} & \cdots & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{\Sigma}_2 & \cdots & \boldsymbol{0} \\ \vdots & \ddots & \ddots & \vdots \\ \boldsymbol{0} & \boldsymbol{0} & \cdots & \boldsymbol{\Sigma}_m \end{pmatrix}.$$

• We assume that $\Sigma_1 \neq \cdots \neq \Sigma_m$.

- If the data are independent (within and between group), but heteroscedastic between groups, the matrix $\Sigma_i = \sigma_i^2 \mathbf{I}$, where \mathbf{I} is the identity matrix with ones on the diagonal and zero for off-diagonal entries.
- In this case, there are *m* variance parameters to estimate (one per group).
- We could use a different structure for Σ_i . SAS allows this, but the blocks cannot share parameters, so we get *m* times the number of parameters in Σ_i . There must be enough observations in each group to reliably estimate the covariance parameters.

The college data set provides the nine-month academic salary (in thousand dollars) in 2008–2009 of professors in a college in the USA.

- salary: nine month income (in thousand dollars).
- rank: academic rank of the professor (assistant , associate or full).
- field: categorical variable indicating whether research field is applied or theoretical.
- sex: sex of individual, either man or woman.



The explanatory data analysis shows clear heteroscedasticity within academic rank.

SAS code for a different variance per group

```
proc mixed data=statmod.college plots=studentpanel;
class field rank sex;
model salary = sex field rank;
repeated / group = rank;
run;
```

The argument repeated / group specifies the group structure.

Cov Parm	Group	Estimate	Null Model Likelihood R	atio	
Residual	rank assistant	42.4817	Test	Test	
Residual	rank associate	115.29	DF Chi-Square Pr > Cl	niSq	
Residual	rank full	722.44	2 164.78 <.	0001	

The variance increases with rank. The likelihood ratio test shows that the model with a different group for each rank is significantly better than the linear model which assumes a constant variance for every observation.

Diagnostic plots for the profsalaries data



The residual plots show that the model captures most features well. We can be confident in our inference.

Type 3 Tests of Fixed Effects							
Effect	Num DF	Den DF	F Value	Pr > F			
sex	1	392	1.55	0.2141			
field	1	392	92.85	<.0001			
rank	2	392	334.46	<.0001			

- Solely comparing the salary of mean and women academics using a two-sample test is wrong, because rank is an important explanatory variable.
- Moreover, the proportion of full professor that are women (7%) is much lower than for assistant or associated professors (16%)
- After accounting for rank and dealing with group heteroscedasticity, there is no evidence of gender gap.