# Effect size and power

#### **Session 8**

MATH 80667A: Experimental Design and Statistical Methods for Quantitative Research in Management HEC Montréal

#### Outline





Effect size

### Motivating example

Quote from the OSC psychology replication

The key statistics provided in the paper to test the "depletion" hypothesis is the main effect of a one-way ANOVA with three experimental conditions and confirmatory information processing as the dependent variable; F(2, 82) = 4.05, p = 0.02,  $\eta^2 = 0.09$ . Considering the original effect size and an alpha of 0.05 the sample size needed to achieve 90% power is  $_{132}$  subjects.

Replication report of Fischer, Greitemeyer, and Frey (2008, JPSP, Study 2) by E.M. Galliani

#### Translating statement into science

Q: How many observations should I gather to reliably detect an effect?

Q: How big is this effect?

#### Does it matter?

#### Statistical significance $\neq$ practical relevance

With large enough sample size, **any** sized difference between treatments becomes statistically significant.

But whether this is important depends on the scientific question.

### Example

- What is the minimum difference between two treatments that would be large enough to justify commercialization of a drug?
- Tradeoff between efficacy of new treatment vs status quo, cost of drug, etc.

### Using statistics to measure effects

Statistics and *p*-values are not good summaries of magnitude of an effect:

• the larger the sample size, the bigger the statistic, the smaller the *p*-value

Instead use

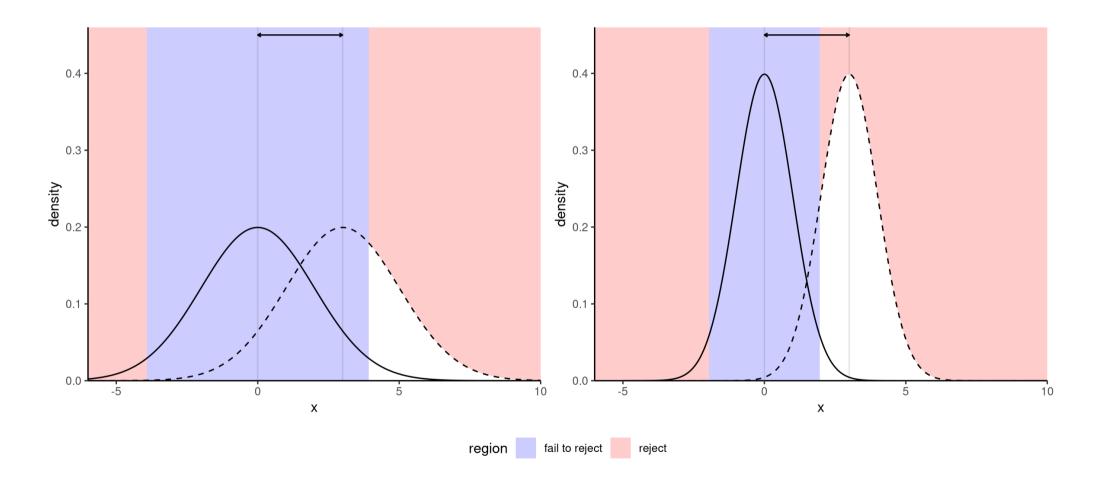
#### standardized differences

percentage of variability explained

Estimators popularized in the handbook

Cohen, Jacob. Statistical Power Analysis for the Behavioral Sciences, 2nd ed., Routhledge, 1988.

#### Illustrating effect size (differences)



The plot shows null (thick) and true sampling distributions (dashed) for sample mean with small (left) and large (right) samples.

#### Estimands, estimators, estimates

- $\mu_i$  is the (unknown) population mean of group *i* (parameter, or estimand)
- $\hat{\mu}_i$  is a formula (an estimator) that takes data as input and returns a numerical value (an estimate).
- throughout, use hats to denote estimated quantities:







Left to right: parameter  $\mu$  (target), estimator  $\widehat{\mu}$  (recipe) and estimate  $\widehat{\mu}=10$  (numerical value, proxy)

### Cohen's d

Standardized measure of effect (dimensionless=no units):

Assuming equal variance  $\sigma^2$ , compare mean of two groups *i* and *j*:

$$d=rac{\mu_i-\mu_j}{\sigma}$$

- Cohen's d's usual estimator d uses sample average of groups and the pooled variance estimator σ.
- Hedge's g includes a finite sample correction and is less biased (prefered in practice)

Cohen's classification of effect sizes: small (*d*=0.2), medium (*d*=0.5) or large (*d*=0.8).

### Cohen's *f*

For a one-way ANOVA (equal variance  $\sigma^2$ ) with more than two groups,

$$f^2=rac{1}{\sigma^2}\sum_{j=1}^krac{n_j}{n}(\mu_j-\mu)^2,$$

a weighted sum of squared difference relative to the overall mean  $\mu$ . For k = 2 groups, Cohen's f and Cohen's d are related via f = d/2.

#### Effect size: proportion of variance

If there is a single experimental factor, use **total** effect size. Break down the variability

$$\sigma_{
m total}^2 = \sigma_{
m resid}^2 + \sigma_{
m effect}^2$$

and define the percentage of variability explained by the effect.

$$\eta^2 = rac{ ext{explained variability}}{ ext{total variability}} = rac{\sigma_{ ext{effect}}^2}{\sigma_{ ext{total}}^2}$$

#### Coefficient of determination estimator

## For the balanced one-way between-subject ANOVA, typical estimator is the **coefficient of determination**

$$\widehat{R}^2 = rac{F 
u_1}{F 
u_1 + 
u_2}$$

where  $\nu_1 = K - 1$  and  $\nu_2 = n - K$  are the degrees of freedom for the one-way ANOVA with *n* observations and *K* groups.

- $\hat{R}^2$  is an upward biased estimator (too large on average).
- People frequently write  $\eta^2$  when they mean  $\hat{R}^2$
- for the replication,  $\hat{R}^2 = (4.05 \times 2)/(4.05 \times 2 + 82) = 0.09$

#### $\omega^2$ square estimator

Another estimator of  $\eta^2$  that is recommended in Keppel & Wickens (2004) for power calculations is  $\hat{\omega}^2$ .

For one-way between-subject ANOVA, the latter is obtained from the *F*-statistic as

$$\widehat{\omega}^2 = rac{
u_1(F-1)}{
u_1(F-1)+n}$$

- for the replication,  $\hat{\omega}^2 = (3.05 \times 2)/(3.05 \times 2 + 84) = 0.0677.$
- if the value returned is negative, report zero.

#### Converting $\eta^2$ to Cohen's f

Software usually take Cohen's  $_f$  (or  $_{f^2}$ ) as input for the effect size. Convert from  $_{\eta}$  to  $_f$  via the relationship

$$f^2=rac{\eta^2}{1-\eta^2}.$$

#### If we plug-in estimated values

- with  $\hat{R}^2$ , we get  $\hat{f} = 0.314$
- with  $\widehat{\omega}^2$ , we get  $\widetilde{f} = 0.27$ .

#### Effect sizes for multiway ANOVA

With a completely randomized design with only experimental factors, use **partial** effect size

$$\eta^2_{
m \langle effect 
angle} = \sigma^2_{
m effect} / (\sigma^2_{
m effect} + \sigma^2_{
m resid})$$

In R, USe effectsize::omega\_squared(model, partial = TRUE).

#### Partial effects and variance decomposition

Consider a completely randomized balanced design with two factors *A*, *B* and their interaction *AB*. We can decompose the total variance as

$$\sigma_{ ext{total}}^2 = \sigma_A^2 + \sigma_B^2 + \sigma_{AB}^2 + \sigma_{ ext{resid}}^2$$

Cohen's partial *f* measures the proportion of variability that is explained by a main effect or an interaction, e.g.,

$$f_{\langle A 
angle} = rac{\sigma_A^2}{\sigma_{ ext{resid}}^2}, \qquad f_{\langle AB 
angle} = rac{\sigma_{AB}^2}{\sigma_{ ext{resid}}^2}.$$

#### Partial effect size (variance)

Effect size are often reported in terms of variability via the ratio

$$\eta^2_{
m \langle effect 
angle} = rac{\sigma^2_{
m effect}}{\sigma^2_{
m effect} + \sigma^2_{
m resid}}$$

• Both  $\hat{\eta}^2_{\langle \text{effect} \rangle}$  (aka  $\hat{R}^2_{\langle \text{effect} \rangle}$ ) and  $\hat{\omega}^2_{\langle \text{effect} \rangle}$  are **estimators** of this quantity and obtained from the *F* statistic and degrees of freedom of the effect.

#### Estimation of partial $\omega^2$

Similar formulae as the one-way case for between-subject experiments, with

 $\widehat{\omega}^2_{\langle ext{effect}
angle} = rac{ ext{df}_{ ext{effect}}(F_{ ext{effect}}-1)}{ ext{df}_{ ext{effect}}(F_{ ext{effect}}-1)+n},$ 

where n is the overall sample size.

In **R**, effectsize::omega\_squared reports these estimates with one-sided confidence intervals.

Reference for confidence intervals: Steiger (2004), Psychological Methods

#### Converting $\omega^2$ to Cohen's f

Given an estimate of  $\eta^2_{(\text{effect})}$ , convert it into an estimate of Cohen's partial  $f^2_{(\text{effect})}$ , e.g.,

$${\widehat{f}}^2_{\langle {
m effect} 
angle} = rac{{\widehat{\omega}}^2_{\langle {
m effect}} 
angle}{1 - {\widehat{\omega}}^2_{\langle {
m effect}} 
angle}.$$

The package effectsize::cohens\_f returns  $\tilde{f}^2 = n^{-1}F_{\text{effect}}$ , a transformation of  $\hat{\eta}^2_{\langle \text{effect} \rangle}$ 

٠

#### Semipartial effect sizes

If there is a mix of experimental and blocking factor...

Include the variance of all blocking factors and interactions (only with the effect!) in denominator.

• e.g., if *A* is effect of interest, *B* is a blocking factor and *C* is another experimental factor, use

$$\eta^2_{\langle A 
angle} = rac{\sigma^2_A}{\sigma^2_A + \sigma^2_B + \sigma^2_{AB} + \sigma^2_{ ext{resid}}}$$

In R, use effectsize::omega\_squared(model, partial = TRUE, generalized = "blocking")
where blocking gets replaced with a vector containing the name of the blocking factors.

### Summary

- Effect sizes can be recovered using information found in the ANOVA table.
- Multiple estimators for the same quantity
  - report the one used along with confidence or tolerance intervals.
- The correct measure may depend on the design
  - partial vs total effects,
  - different formulas for within-subjects (repeated measures) designs!
- Include blocking as part of the variability considered.



#### Power and sample size calculations

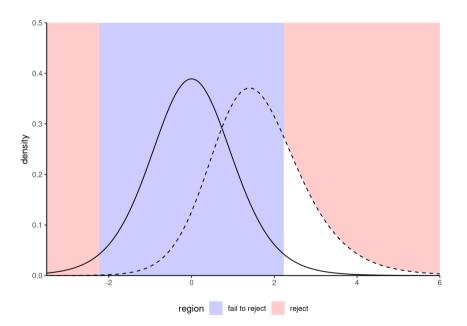
Journals and grant agencies oftentimes require an estimate of the sample size needed for a study.

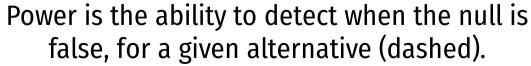
- large enough to pick-up effects of scientific interest (good signal-to-noise)
- efficient allocation of resources (don't waste time/money)

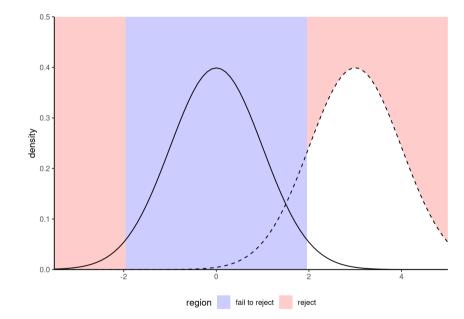
Same for replication studies: how many participants needed?

#### I cried power!

The null alternative corresponds to a single value (equality in mean), whereas there are infinitely many alternatives...







Power is the area in white under the dashed curved, beyond the cutoff.

#### What determines power?

Think in your head of potential factors impact power for a factorial design.

The size of the effects, δ<sub>1</sub> = μ<sub>1</sub> - μ, ..., δ<sub>K</sub> = μ<sub>K</sub> - μ
 The background noise (intrinsic variability, σ<sup>2</sup>)
 The level of the test, α
 The sample size in each group, n<sub>j</sub>
 The choice of experimental design
 The choice of test statistic

We focus on the interplay between

effect size | power | sample size

#### Living in an alternative world

In a one-way ANOVA, the alternative distribution of the  $_F$  test has an additional parameter  $_\Delta$ , which depends on both the sample and the effect sizes.

$$\Delta = rac{\sum_{j=1}^K n_j (\mu_j - \mu)^2}{\sigma^2} = n f^2$$

Under the null hypothesis,  $\mu_j = \mu$  for j = 1, ..., K and  $\Delta = 0$ .

The greater  $\triangle$ , the further the mode (peak of the distribution) is from unity.

#### Noncentrality parameter and power

$$\Delta = rac{\sum_{j=1}^K n_j (\mu_j - \mu)^2}{\sigma^2}.$$

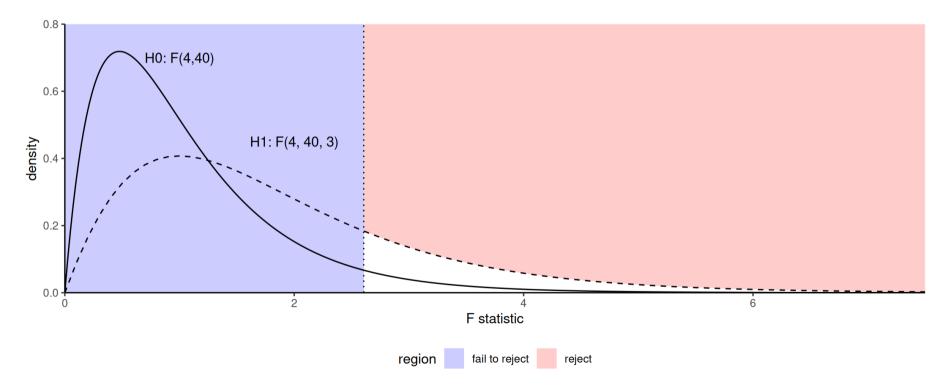
#### When does power increase?

What is the effect of an increase of the

- group sample size  $n_1, \ldots, n_K$ .
- variability  $\sigma^2$ .
- true mean difference  $\mu_j \mu_i$ .

#### Noncentrality parameter

The alternative distribution is  $_{F(\nu_1,\nu_2,\Delta)}$  distribution with degrees of freedom  $_{\nu_1}$  and  $_{\nu_2}$  and noncentrality parameter  $_{\Delta}$ .



#### Power for factorial experiments

- $G^*Power$  and **R** packages take Cohen's  $f(or f^2)$  as inputs.
- Calculation based on *F* distribution with
  - $\circ \quad {\it $\nu_1={\rm df}_{\rm effect}$} \ degrees \ of \ freedom$
  - $\circ \nu_2 = n n_g$ , where  $n_g$  is the number of mean parameters estimated.
  - $\circ$  noncentrality parameter  $\phi = n f_{\langle ext{effect} 
    angle}^2$

### Example

#### Consider a completely randomized design with two crossed factors $_A$ and $_B$ . We are interested by the interaction, $\eta^2_{(AB)}$ , and we want 80% power:

#### Power curves

```
library(pwr)
power_curve <-
pwr.anova.test(
  f = 0.314, #from R-squared
  k = 3,
  power = 0.9,
  sig.level = 0.05)
plot(power_curve)</pre>
```

Recall: convert  $\eta^2$  to Cohen's f (the effect size reported in pwr) via  $f^2 = \eta^2/(1-\eta^2)$ Using  $\tilde{f}$  instead (from  $\hat{\omega}^2$ ) yields n = 59 observations per group!

#### Balanced one-way analysis of variance power calculation 100% groups k = 3effect size f = 0.314 alpha = 0.0575% power 50% 25% optimal sample size n = 44n is number in each group 0% 20 60 sample size per group

#### Effect size estimates

### WARNING!

Most effects reported in the literature are severely inflated.

Publication bias & the file drawer problem

- Estimates reported in meta-analysis, etc. are not reliable.
- Run pilot study, provide educated guesses.
- Estimated effects size are uncertain (report confidence intervals).

#### Beware of small samples

Better to do a large replication than multiple small studies. Otherwise, you risk being in this situation:



### Observed (post-hoc) power

Sometimes, the estimated values of the effect size, etc. are used as plug-in.

- The (estimated) effect size in studies are noisy!
- The post-hoc power estimate is also noisy and typically overoptimistic.
- Not recommended, but useful pointer if the observed difference seems important (large), but there isn't enough evidence (too low signal-to-noise).

#### **Statistical fallacy**

Because we reject a null doesn't mean the alternative is true!